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H. Rose Vineer, S.H. Verschave, E. Claerebout, J. Vercruysse, D.J. Shaw, J. Charlier, E.R. Morgan

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**GLOWORM-PARA: a flexible framework to simulate the population  
dynamics of the parasitic phase of gastrointestinal nematodes infecting  
grazing livestock**

Rose Vineer, H.<sup>a,b,c,\*</sup>, Verschave, S. H.<sup>d,e</sup>, Claerebout, E.<sup>d</sup>, Vercruysse, J.<sup>d</sup>, Shaw, D.J.<sup>f</sup>,  
Charlier, J.<sup>g</sup>, Morgan, E. R.<sup>a,b,h</sup>

<sup>a</sup> *Veterinary Parasitology and Ecology Group, Bristol Veterinary School, University of  
Bristol, UK, BS8 1TQ*

<sup>b</sup> *Cabot Institute, Royal Fort House, University of Bristol, UK, BS8 1UJ*

<sup>c</sup> *Department of Infection Biology, Institute of Infection and Global Health, University  
of Liverpool, Leahurst Campus, Neston, Cheshire, UK, CH64 7TE*

<sup>d</sup> *Department of Virology, Parasitology and Immunology, Faculty of Veterinary  
Medicine, Ghent University, Salisburylaan 133, 9820 Merelbeke, Belgium*

<sup>e</sup> *Department of Molecular and Cellular Biology, Harvard University, 52 Oxford Street,  
Cambridge, MA 02138, USA*

<sup>f</sup> *The Royal (Dick) School of Veterinary Studies and The Roslin Institute, University of  
Edinburgh, Easter Bush Campus, Roslin, EH25 9RG, United Kingdom*

<sup>g</sup> *Kreavet, Hendrik Mertensstraat 17, 9150 Kruibeke, Belgium*

<sup>h</sup> *Institute for Global Food Security, Queen's University Belfast, UK, BT9 7BL*

\*Corresponding author. E-mail address: [hannah.vineer@liverpool.ac.uk](mailto:hannah.vineer@liverpool.ac.uk)

24 **ABSTRACT**

25 Gastrointestinal (GI) nematodes are a significant threat to the economic and  
26 environmental sustainability of keeping livestock, as adequate control becomes  
27 increasingly difficult due to the development of anthelmintic resistance (AR) in some  
28 systems and climate-driven changes to infection dynamics. To mitigate any negative  
29 impacts of climate on GI nematode epidemiology and slow AR development, there is  
30 a need to develop effective, targeted control strategies that minimise the  
31 unnecessary use of anthelmintic drugs and incorporate alternative strategies such as  
32 vaccination and evasive grazing. However, the impacts climate and GI nematode  
33 epidemiology may have on the optimal control strategy are generally not  
34 considered, due to lack of available evidence to drive recommendations. Parasite  
35 transmission models can support control strategy evaluation to target field trials,  
36 thus reducing the resources and lead-time required to develop evidence-based  
37 control recommendations incorporating climate stochasticity. GI nematode  
38 population dynamics arising from natural infections have been difficult to replicate  
39 and model applications have often focussed on the free-living stages. A flexible  
40 framework is presented for the parasitic phase of GI nematodes, GLOWORM-PARA,  
41 which complements an existing model of the free-living stages, GLOWORM-FL.  
42 Longitudinal parasitological data for two species that are of major economic  
43 importance in cattle, *Ostertagia ostertagi* and *Cooperia oncophora*, were obtained  
44 from seven cattle farms in Belgium for model validation. The framework replicated  
45 the observed seasonal dynamics of infection in cattle on these farms and overall,  
46 there was no evidence of systematic under- or over-prediction of faecal egg counts  
47 (FECs). However, the model under-predicted the FECs observed on one farm with

48 very young calves, highlighting potential areas of uncertainty that may need further  
49 investigation if the model is to be applied to young livestock. The model could be  
50 used to drive further research into alternative parasite control strategies such as  
51 vaccine development and novel treatment approaches, and to understand GI  
52 nematode epidemiology under changing climate and host management.

53

54 *Keywords: Ostertagia ostertagi; Cooperia oncophora; Model; Parasite; Population*  
55 *dynamics; Transmission; Nematode; Livestock*

56

## 1. Introduction

Gastrointestinal (GI) nematodes are increasingly recognised as an important threat to the future sustainability of keeping livestock for food production and leisure. At a policy level, livestock make a significant contribution to agricultural greenhouse gas (GHG) emissions, which may be exacerbated by GI nematode infections (Fox et al., 2018). In 2013, methane emissions from cattle and sheep were responsible for 47% of agricultural GHG emissions in England, and approaching 90% of agricultural GHG emissions in Wales, Scotland and Northern Ireland (Salisbury et al., 2015). GI nematodes also threaten food security and the economic sustainability of livestock farming as they cause significant production losses in ruminants (Nieuwhof and Bishop, 2005; Charlier et al., 2009). For example, a meta-analysis of 88 studies found that lambs infected with GI nematodes experienced a 25% reduction in weight gain (Mavrot et al., 2015). Similarly, in cattle, GI nematodes cause significant reductions in weight gain and milk yield (Verschave et al., 2014b). Finally, the pathogenic implications of GI nematode infections (e.g. Besier et al., 2016) on host welfare are clear, however the impact of subclinical and chronic infections remains an understudied but important question in livestock helminth research (Morgan et al., 2018).

Currently, the control of GI nematodes in livestock is primarily based on the chemotherapeutic use of anthelmintic substances (Charlier et al., 2014). However, both the influence of climate change and, in some places, the development of AR on farm management and parasite epidemiology are expected to challenge the future control of these infections (Morgan and van Dijk, 2012; Skuce et al., 2013). Progress has been made towards targeted, sustainable control strategies that are

economically sound (Charlier et al., 2014) but the need for adequate decision-support tools to aid in the implementation of these strategies remains (Morgan, 2013). Multiple initiatives promoting the sustainable control of parasites in livestock have been developed worldwide, such as SCOPS (Sustainable Control of Parasites in Sheep; [scops.org.uk](http://scops.org.uk)), COWS (Control of Worms Sustainably; [cattleparasites.org.uk](http://cattleparasites.org.uk)), the UK-VET guidelines on parasite control in horses (Rendle et al., 2019) in the UK, and ASKBILL (Kahn et al., 2017) in Australia. These initiatives provide flexibility to adapt their guidelines to different livestock management systems and, in the case of SCOPS, resource-intensive longitudinal studies evaluating the efficacy of their guidelines in a range of systems are ongoing (e.g. Learmount et al., 2018).

However, the epidemiology of GI nematode infections is a result of complex interactions between parasite, host, climate, farm management and historic control strategies. The efficacy of these guidelines in the diversity of management systems practiced by cattle, sheep, goat and horse keepers worldwide, and their sustainability in the face of climate change and AR cannot be studied empirically without significant resource input and multi-year studies to incorporate inter-annual weather variability and extreme weather events.

Parasite transmission models are useful as they provide the potential to include a variety of processes on different levels and extrapolate current knowledge to alternative scenarios at large temporal scales (Rose et al., 2015). In doing so, model simulations can be used to target resources for empirical research where they are needed most, and guide the development of evidence-based parasite control strategies and tools. The development of mathematical models to simulate the transmission dynamics of GI nematode infections in ruminants dates back several

decades. The majority of the existing models were developed specifically for GI nematode infections in sheep (Verschave et al., 2016a lists 32 models). Fewer models exist for cattle nematodes (Verschave et al., 2016a lists seven models for *Ostertagia ostertagi*), and model development for equine nematode species dates back only a few years (Leathwick et al., 2015, 2016, 2017). Generic models that provide a framework for GI nematode infections that can be applied to a range of host and nematode species are also scarce (Smith, 2011), while their development is of great interest in identifying emerging patterns of change (Molnár et al., 2013; Rose et al., 2014).

Recently, a generic model framework for the free-living phase of GI nematodes, which has important modifications on behaviour and development of the larvae on pasture, was developed (GLOWORM-FL, Rose et al., 2015). This model was initially applied to three species of importance in cattle (*O. ostertagi*) and small ruminants (*Haemonchus contortus* and *Teladorsagia circumcincta*), and additional published parameters and data are available to adapt the model to equine cyathostomins (Leathwick et al., 2015), *Cooperia oncophora* in cattle (Sauermann and Leathwick, 2018) and *Marshallagia marshalli* in small ruminants (Carlsson et al., 2013). To fully explore the consequences of different control and management approaches on parasite epidemiology, however, a complementary model for the parasitic phase is needed as host-parasite interactions and host acquired immunity play crucial parts in the transmission dynamics of GI nematodes (Claerebout and Vercruysse, 2000).

The aim of the current study was to develop a conceptual model framework for the parasitic phase of GI nematodes, GLOWORM-PARA, that can be applied to a range of host and parasite species. Here, the model is parameterised and validated



for two species that are of major economic importance in cattle, i.e. *O. ostertagi* and *C. oncophora*. Previous cattle models have tended to only focus on one nematode species, i.e. *O. ostertagi*, perhaps due to its pathogenic significance compared with *C. oncophora*, against which cattle develop an effective immune response. No single-species model exists for *C. oncophora*, despite its increasing importance in the context of anthelmintic resistance and treatment failure (Sutherland and Leathwick, 2011). The development of acquired immunity against GI nematodes was modelled and parameterised in a heuristic but data-driven manner, to provide a transparent and replicable approach. An extensive set of field observations of first season grazing cattle was used for model validation.

## 2. Materials and methods

### 2.1. Model framework

The model framework (Fig. 1), tracks the mean number of GI nematodes and level of acquired immunity in a group of hosts (i.e. is a population-based mean-field model). State variables and model parameters are defined in Table 1.

Infective L3 are ingested with herbage ( $L3i$ ) and enter the pool of pre-adult parasitic nematodes ( $P$ ). Pre-adult nematodes either develop to adult nematodes ( $A$ ) or arrest their development as larvae ( $Pa$ ) before developing to the adult stage. Although the model is developed and validated for trichostrongylid nematodes in the present study, it could also be applied to other strongylid species with a broadly similar life cycle (e.g. equine cyathostomins) as all pre-adult stages are modelled as a single state variable and the pre-adult stage involved in arrested development is not specified. This basic representation of the GI nematode life cycle is similar to

153 numerous previous models, as reviewed by Verschave et al. (2016a) and Smith  
 154 (2011).

$$\frac{dP}{dt} = L3i - \delta P - \mu_1 P \quad (1)$$

$$\frac{dPa}{dt} = -h_2 \mu_2 Pa + \delta h_1 P \quad (2)$$

$$\frac{dA}{dt} = \delta(1 - h_1)P + h_2 Pa - \mu_3 A \quad (3)$$

155 Acquired immunity ( $r$ ) increases in response to exposure to infection (Claerebout  
 156 and Vercruysse, 2000), in this case the L3 ingestion rate ( $L3i$ ), and decays with time,  
 157 similar to previous models e.g. Roberts and Grenfell (1991). However, the present  
 158 framework differs from previous models in its representation of  $r$  as a logistic growth  
 159 function to facilitate modelling interactions between the host immune response and  
 160 parasite life-history parameters (Section 2.3.3).

$$\frac{dr}{dt} = \rho(L3i)(1 - r) - \sigma r \quad (4)$$

161

## 162 2.2. Model integration

163 The model was implemented in R v 3.5.1 “Feather Spray” (R Core Team, 2018. R:  
 164 A language and environment for statistical computing. R Foundation for Statistical  
 165 Computing, Vienna, Austria) using the *lsoda* function of the *deSolve* package v 1.24  
 166 (Soetaert et al., 2010) for solving differential equations. The model returns daily  
 167 output. Anthelmintic treatments (if applicable) are implemented using the *events*  
 168 argument of the *lsoda* function to reduce the worm burden by a representative  
 169 percentage for the duration of the residual activity of the product used. For example,

an effective moxidectin pour-on treatment for cattle would trigger a reduction in total worm burden of 99% for 35 days. The percentage of reduction could be modified, if necessary, to represent vaccination strategies or reduced anthelmintic efficacies observed in the field. Model output is the mean stage-specific worm burden and egg output per host, which can be used to estimate the faecal egg count (FEC; eggs per gram of faeces (epg)) if the amount of faeces produced is known.

### 2.3. Parameter estimates

Parameterisation of the framework for multiple species is demonstrated using two species infecting cattle that are currently the focus of vaccine development programmes (Matthews et al., 2016): the abomasal nematode *O. ostertagi* and the intestinal nematode *C. oncophora* (Table 2).

#### 2.3.1. Seasonally variable parameters

*Arrested development.* There is currently no consensus on the mechanisms of arrested (hypobiosis) and subsequent resumed development in trichostrongylid nematodes, and numerous confounding factors in available data prevent the development of robust mechanistic models of arrest (Smith, 1974; Michel et al., 1976; Frank et al., 1986; 1988; Eysker, 1993; Fernández et al., 1999; Langrova and Jankovska, 2004; Lützelshwab et al., 2005; Langrova et al., 2008). As the numerous potential drivers of arrest are correlated and seasonal (e.g. the age structure of host populations, host immunity, temperature, moisture and photoperiod), a simplified seasonal approach was taken to estimate seasonal variations in the factor driving

arrest rates ( $d$ ) which is used to simulate the proportion of developing pre-adult nematodes that enter a state of arrested development ( $h_1$ ).

For *C. oncophora* and *O. ostertagi* this arrest rate ( $h_1$ ) was assumed to be related to the development success of eggs and larvae on pasture.  $d$  was approximated as a 7 day moving average, determined by the temperature-dependent development rate (using the *na.ma* function in the *imputeTS* R package v 3.0 (Moritz and Bartz-Beielstein, 2017) and the *ma* function in the *forecast* package v 8.9 (Hyndman and Khandakar, 2008)), whereby the minimum arrest rate is assumed to coincide with the period where development success is at its maximum and the maximum arrest rate is assumed to coincide with the period where development success is at its minimum. Thus, the arrest rate at the current time point,  $t$ , is a function of the species-specific minimum and maximum arrest rates ( $h_{min}$  and  $h_{max}$ ), annual minimum and maximum predicted development success at the study site ( $d_{min}$  and  $d_{max}$ ), and predicted development success at the current timestep ( $d_t$ ).

$$h_1 = h_{(max)} - \left( \frac{h_{(max)} - h_{(min)}}{d_{(max)}} \right) \times d_t \quad (5)$$

The temperature-dependent development rate of eggs and larvae on pasture for *O. ostertagi* was as described in Rose et al. (2015). For *C. oncophora*, development rate data presented in Sauermann and Leathwick (2018) were extracted using Plot Digitizer v2.6.8 (<http://plotdigitizer.sourceforge.net/>), and a linear model applied to the data using the *lm()* function in R.

The proportion of arrested larvae resuming development ( $h_2$ ) was assumed to be an inverse function of the driver of arrest (i.e. development success for *O. ostertagi* and *C. oncophora*):

$$h_2 = \left( \frac{1}{d_{(max)} - d_{(min)}} \right) \times (d_t - d_{(min)}) \quad (6)$$

*L3 ingestion rate and dry matter intake.* To calculate the L3 ingestion rate ( $L3i$ ) the average daily dry matter intake (DMI) by grazing cattle was estimated using the equations of MAFF (1975) based on bodyweight (estimated from the bodyweight at turn out using standard age-related growth curves for dairy cattle; Cue et al., 2012, Verschave et al., 2014a). The equations for growing young stock and adult cattle were used for animals with a bodyweight of less than and more than 400 kg, respectively. The rate of ingestion of dry matter was estimated as a proportion of the total available herbage consumed ( $kgDM$ ; standing biomass, i.e. kilograms of dry matter per hectare). From this, the L3 ingestion rate was estimated as follows (parameter and state variables are defined in Table 1):

$$L3i = -\ln\left(1 - \frac{DMI}{kgDM}\right) \times L3h \quad (7)$$

*Faeces production and faecal egg counts.* The average daily faeces production was estimated based on host body weight using the formula of Nennich et al. (2005). Mean FECs (epg) for the group of hosts can be estimated from the number of adults ( $A$ ), per adult fecundity rate ( $\lambda$ ) and expected daily faeces production ( $f$ ).

$$\text{FEC} = \frac{Ae^{\lambda}}{f} \quad (8)$$

234

235 *2.3.2. Constant rates*

236 The development rate ( $\delta$ ) from ingested L3 to mature adult was estimated from a  
 237 prepatent period of 17 days for both *O. ostertagi* and *C. oncophora*; (Table 2; Powers  
 238 et al., 1982).

239 No data were available in the literature to estimate the mortality rates of arrested  
 240 larvae due to the confounding effects of resumed development. Therefore, the  
 241 mortality rate of arrested L4s for both *O. ostertagi* and *C. oncophora* was set at 0.002  
 242 after Grenfell et al. (1987). Mortality rates of all other pre-adult and adult  
 243 nematodes were a function of immunity (section 2.3.3).

244

245 *2.3.3. Immunity and dependent parameters*

246 *Immune response and decay rate.* No data were available to formally estimate  
 247 immunity decay rates ( $\sigma$ ) for *O. ostertagi* nor *C. oncophora*, therefore three experts  
 248 in the area of cattle GI nematodes and vaccine development (J. Vercruysse, E.  
 249 Claerebout (both co-authors on this study) and P. Dorny, Ghent University, Belgium)  
 250 were consulted in order to estimate the percentage of decay in immunity over a  
 251 typical housing period.

252 To estimate the response rate ( $\rho$ ), it was assumed that protective immunity ( $r=1$ )  
 253 was typically acquired after 1.5 grazing seasons (9 months of exposure punctuated  
 254 by a 6 month housing period during which immunity is assumed to decay as  
 255 described above) for *O. ostertagi* and one grazing season (6 months) for *C.*

*oncophora* (Armour, 1989; Ploeger et al., 1995; Claerebout et al., 1998; Ravinet et al., 2014). Species-specific field observations of L3 density on pasture ( $L3h$ ) over the course of a grazing season were extracted from the raw data from field trials across Europe summarised by Shaw et al. (1998). The data concerned pasture larval counts from both 'clinical' and 'subclinical' pastures (i.e. pastures on which an outbreak of parasitic gastroenteritis in the untreated first season grazers was observed, or not, respectively) and included a mixture of calves that were treated with anthelmintics and untreated controls (Shaw et al., 1998). Using these data, equation 4, the decay rate ( $\sigma$ ) and the method described in section 2.3.1 for estimating L3 intake rates, the *optimise* function in R was used to find the response rate that minimised the sum of square error (SSE) for each dataset, given the expectation that  $r$  should equal 0.4 and 0.6 after 3 months of grazing, and 0.7 and 1 after 6 months of grazing for *O. ostertagi* and *C. oncophora*, respectively.

*Immunity-mediated regulation of the parasite population.* Host acquired immunity is assumed to regulate the parasite population in three ways: (i) by exclusion of ingested larvae (increased pre-adult mortality rate); (ii) by decreasing the survival of established (adult) nematodes; and (iii) by decreasing the fecundity of adult nematodes (Barger et al., 1985; Smith and Grenfell, 1985; Coyne and Smith, 1992; Smith, 1994; Stear et al., 1995; Claerebout and Vercruysse, 2000; Garnier et al., 2016). Thus immunity-mediated regulation of the parasite population was incorporated by increasing the mortality rates of pre-adult ( $\mu_1$ ) and adult nematodes ( $\mu_3$ ), and decreasing fecundity ( $\lambda$ ) with increasing acquired immunity. As acquired immunity cannot be measured directly (Claerebout and Vercruysse, 2000), little is

known about the functional relationship between acquired immunity and these parameters. Therefore, a linear relationship was assumed, whereby mortality increases between the minimum and maximum values, and fecundity decreases between the maximum and minimum values as acquired immunity increases between 0 and 1:

$$\mu_i = \mu_{i(min)} + (\mu_{i(max)} - \mu_{i(min)})r \quad (9)$$

$$\lambda = \lambda_{(max)} - (\lambda_{(max)} - \lambda_{(min)})r \quad (10)$$

## 2.4. Model validation

### 2.4.1. Longitudinal data

The model was validated using independent datasets containing longitudinal parasitological data collected during 2012 and 2013 (described in detail in Verschave, S.H., 2015. Development of a transmission model for gastro-intestinal nematode infections in cattle. PhD thesis. Ghent University, Belgium) and summarised in Supplementary Table S1. The sampled herds consisted of first season grazers located on seven commercial dairy farms in Flanders (Belgium). The herds were visited monthly from turn out in Spring (April, May or June) until housing in Autumn (September, October or November).

FECs of all animals were performed each month using a modified McMaster technique with a sensitivity of 10 epg (MAFF, 1986) and the mean and 95% confidence interval estimated for each month. For this, the *sample* and *replicate*



functions in R were used to generate 10,000 replicates sampling with replacement. The *quantiles* function was then used with probabilities of 0.025 and 0.975 to obtain the bootstrapped 95% confidence limits for the means of these replicates.

For nematode species identification, the positive faecal samples were mixed per herd, cultured and identified according to Borgsteede and Hendriks (1973). Pasture infectivity (density of L3 on herbage; *L3h*) was measured as described in Verschave et al. (2015) each month and every 2 months, respectively, in 2012 and 2013 using the modified technique of Taylor (1939).

#### 2.4.2. Validation simulations

Mean FECs and 95% confidence intervals reported by Verschave (Verschave, S.H., 2015. PhD thesis, cited earlier) were corrected for incomplete egg recovery (recovery rate of 55%; Paras et al., 2018). Corresponding daily pasture contamination values for the entire period of each trial were obtained by interpolation of the monthly pasture contamination values using the *approxfun* function in R. The data collected from each herd formed a separate validation dataset.

Daily mean air temperature data used to estimate daily values for development success to estimate arrest rates (equations 5 and 6) were obtained for each herd from the E-OBS gridded dataset (Haylock et al., 2008; dataset available to download from <https://www.ecad.eu/download/ensembles/download.php>) based on the location of the village where each herd was located (Supplementary Table S1) using the *ncdf4* function v 1.17 in R (Pierce, D. 2019. *ncdf4*: Interface to Unidata netCDF (Version 4 or Earlier) Format Data Files. R package version 1.17. <https://CRAN.R-project.org/package=ncdf4>).

The longitudinal field observations were used to validate species-specific deterministic model simulations. Daily L3 intake rates were estimated from the interpolated field data as described in equation 7. Dry matter intake and faeces production were estimated as described in section 2.3.1. No data were available for standing biomass, therefore, 2000 kg of DM per hectare was assumed. Although the individuals in the longitudinal datasets were first season grazers (i.e. had never had exposure to *O. ostertagi* nor *C. oncophora*), there is potential for age-related immunity due to the maturation of the immune system (see discussion in Vercruysse and Claerebout, 1997, and Smith et al., 1985). Therefore, host immunity ( $r$ ) was set at an initial value between 0.1 and 0.5, dependent on age at the start of the grazing period (i.e. 6 months of age,  $r = 0.1$ ; 21 months of age,  $r = 0.5$ ).

#### 2.4.3. Statistical validation: deterministic simulations

Model goodness of fit was assessed using a linear regression through the origin of observed and predicted FECs as described by Rose et al. (2015). The first FEC for each herd was excluded from statistical validation as this FEC was simply to confirm the absence of infection at turn out. A statistically significant regression with low residual error indicates that the model reproduces the seasonal patterns of the observed FECs. However, statistical significance does not validate the ability of the model to reproduce the magnitude of FECs observed. For this, the slope estimate was used. A perfect linear fit between model predictions and field observations implies an intercept of zero and a slope of 1. A regression through the origin with a slope that is not significantly different from 1, and therefore included in the 95% confidence interval, indicates that the model reproduces the magnitude of observed

FECs over the course of the season, with values significantly less than 1 indicating underestimation of FECs and values significantly greater than 1 indicating overestimation of FECs. A high  $R^2$  value indicates that the model captures a significant proportion of the variance in the observed FECs. Due to the relatively small number of individuals in each herd, the potential for considerable individual variation in FECs (Levecke et al., 2011), and the limitations of the McMaster's faecal egg counting method and other flotation techniques (Egwang and Slocombe, 1981), visual comparison of observed and predicted values were incorporated into the evaluation to mitigate against this variability undermining statistical validation.

#### 2.4.4. Qualitative validation: stochastic simulations

Although the framework presented here is a deterministic mean-field model representing a group of hosts, incorporating individual variation is possible and is beneficial for further validation and future evaluation of individual-based parasite control strategies. An additional 50 simulations were run per herd, per nematode species (representing 50 individual hosts) to incorporate the stochastic influences of between-host variation. The aggregation of  $L3h$  and chance encounters with  $L3h$  during grazing was incorporated as described in Berk et al. (2016b). The  $L3$  ingestion rate was drawn from a negative binomial distribution using the *rnbinom* function in R, with a mean equal to the observed  $L3h$  at each time point, and a high level of aggregation, as would be expected for the moderate  $L3h$  densities observed in this study ( $k = 1.41$ ; Verschave et al., 2015). For other species or farming systems, the mean and aggregation values could be adapted to reflect the characteristics of the system to be modelled. In addition to stochastic  $L3$  ingestion, between-host

variability in immune response was drawn from a negative binomial distribution with a mean equal to  $\rho$  and level of aggregation equal to that used for the L3 intake rate, after Stear et al. (2007) suggested that the distribution of the immune response between hosts mirrored that of the parasitological variables.

The practical significance of deviations in model predictions from the observed FECs was also considered in the context of the hypothetical use of the simulated FECs to guide further risk assessment (e.g. prompting a FEC or weighing) and potentially trigger anthelmintic treatment in cattle. Fifty to 200 epg is considered a “Medium” to “High” risk egg count (COWS, 2014). Therefore, a deviation of 200 epg in predicted FECs within the range of 0-400 epg could theoretically result in incorrect risk assessment and anthelmintic treatment choices.

### 3. Results

#### 3.1. Parameter estimates

Linear regression of the development rates reported by Sauermann and Leathwick (2018) against temperature for *C. oncophora* yielded a statistically significant fit ( $a = -0.04$ ,  $b = 0.008$ ,  $R^2 = 0.8166$ ,  $R^2_{\text{adjusted}} = 0.8058$ ,  $F_{(1,17)} = 75.7$ ,  $p = 1.142e-07$ ) with a predicted minimum development threshold of 5°C (minimum threshold for development =  $(0-a)/b$ ).

Expert opinion placed the estimated decay rate over an average 6 month housing period (Charlier et al., 2010) at between 10% and 50%. Therefore, a 6 month decay rate of 30% was used to estimate a daily decay rate (Table 2).

Using optimisation to fit the response rate ( $\rho$ ) to data from Shaw et al. (1998) yielded a median response rate for *O. ostertagi* of 0.00006 (inter-quartile range (IQR)

0.00024) with a median SSE of 0.03205 (IQR 0.32502). For *C. oncophora* this yielded a median response rate of 0.00013 (IQR 0.00040) with a median SSE of 0.00250 (IQR 0.11904). The median fitted response rate was used in all subsequent simulations.

All other parameter estimates are provided in Table 2.

### 3.2. Model validation

The R code used for model simulations and validation is provided in Supplementary Data S1. The code can be viewed and run in R software, or viewed in a plain text editor.

Daily temperature and rainfall data are shown for the location of each herd in Supplementary Fig. S1. The average age at turn out varied between 6 and 21 months (Verschave, S.H., 2015. PhD Thesis, cited earlier; Supplementary Table S1). With the exception of herd 2, longitudinal FEC data used for model validation (Figs. 2 and 3; Supplementary Data S1) tended to be low throughout the grazing season. Mean pasture larval counts (L3h kg DM<sup>-1</sup>) were low at turnout, ranging from 0.001 to 176 L3h kg DM<sup>-1</sup> (Supplementary Table S1), potentially accounting for the low FECs. However the FECs used for validation are typical for calves in their first grazing season with subclinical infections (Shaw et al., 1997).

Qualitatively, species-specific simulations for *O. ostertagi* and *C. oncophora* reproduced general observed patterns of FECs over the course of a grazing season in first season grazers (Figs. 2 and 3). Overall, the model captured a high proportion of variance in the observed FECs for both *O. ostertagi* (mean  $R^2 = 0.76$ ) and *C. oncophora* (mean  $R^2 = 0.67$ ), and residual error was low (Table 3). A statistically

significant regression through the origin was achieved for 6/7 (*O. ostertagi*) and 3/7 (*C. oncophora*) of the validation datasets (Table 3; Fig. 4).

For *O. ostertagi*, there were both negative and positive deviations in the slope from 1 (Table 3), indicating under- or over-prediction of FECs, respectively. However, as the model both under- and over-predicted FECs, there was no evidence of systematic bias. Significant deviations in the slope from 1 were estimated for three of the herds (Table 3). Qualitative assessment of model fit against the mean model and 50 individual simulations incorporating individual variation in immune response and the aggregation of L3s on pasture (Fig. 2) suggests that these deviations are of no practical significance (section 2.4.4), with the exception of Herd 2, where high FECs were observed at the end of the grazing season while predicted FECs remained low.

For *C. oncophora*, there were predominantly negative deviations in the slope from 1, indicating underprediction of FECs. A significant deviation in the slope from 1 was estimated for five of the herds (Table 3). Nevertheless, qualitative assessment as above (Fig. 3) suggests that these deviations, similar to the *O. ostertagi* simulations, are of little practical significance (section 2.4.4). Herd 2 was, again, an exception, with higher FECs observed than predicted.

#### 4. Discussion

Smith noted, in 2011, that *“Although it was eventually realised that within each class of parasites a single generic model framework with suitably adjusted parameter values could satisfactorily represent almost all the infections of interest... most of the examples of nematode and trematode models in the literature were constructed on*

an ad hoc basis to address issues dealing with control of a specific parasite in a specific host in a specific country". Although many of the models published in recent decades contain important differences in the focus of detail necessary for the specific application of the model (e.g. Cornell et al., 2004; Laurenson et al., 2011), more widely applicable models use proprietary software (e.g. Learmount et al., 2006) or are developed using complex spreadsheets (e.g. Sauermann and Leathwick 2018) that can be difficult to reproduce. Furthermore, differences in the structure of the various model frameworks available and their parameters prevent direct comparisons between model outputs.

Here, GLOWORM-PARA, a generic model framework for the parasitic phase of GI nematode infections is presented. The model can be adapted to different host and nematode species and was developed to complement a previously published model of the free-living stages (GLOWORM-FL; Rose et al., 2015). The model's flexibility is demonstrated by parameterisation for two economically important trichostrongylid species infecting cattle worldwide, *C. oncophora* and *O. ostertagi*, and validation against multiple independent datasets. To our knowledge, no previous attempt has been made to model *C. oncophora* transmission alone.

Aspects of the framework that can be adapted to represent the host-parasite system of interest are highlighted throughout. For example, parameter  $d$ , the factor driving arrested development, could be adapted to include immunity (e.g. to simulate the periparturient rise in FECs in ewes), or to extend the estimate of development success (used here for *C. oncophora* and *O. ostertagi*) to include the impacts of desiccation to simulate seasonal arrest in arid regions. In addition, reduced weight gain and parasite-induced anorexia are economically important impacts of GI

nematode infection that have been the focus of many previous models (e.g. Laurenson et al., 2011; Berk et al., 2016a) but were not considered here.

Reduced weight gain could be included in the model by substituting the Cue et al. (2012) equation used here with one that estimates weight gain based on worm burden, and estimating DMI as a function of worm burden to simulate parasite-induced anorexia (e.g. Berk et al., 2016a). The simulations presented here also assumed constant herbage biomass throughout the grazing season due to the lack of data and models to track grass growth. Incorporating models of grass growth, and thus seasonal changes in biomass, may improve predictions by acting on the L3 ingestion rate (i.e. rate of infection, which is a function of DMI rates, total L3 on herbage and available standing biomass). Berk et al. (2016b) incorporated mean monthly grass growth rates for England in their model of *O. ostertagi* in calves. However this, and infection-dependent host growth, was beyond the scope of the current study, which was to develop a minimal, location-independent framework that could be easily parameterised for multiple species and host systems.

GLOWORM-PARA is a mean-field model, simulating the mean trajectory of parasite population dynamics and host immunity in a group of hosts. Mean-field models are useful to explore changes in the system under disparate conditions such as current climate and predicted climate change (Rose et al., 2016), and to evaluate the impact of competing management strategies at a herd/flock level (Learmount et al., 2012). However, in an attempt to stem the development of AR, the focus of parasite control has turned from whole-group treatments to targeted selective treatment (TST), where individuals in a flock/herd are treated either based on parasitological indicators (e.g. FECs in horses and sheep; Kenyon and Jackson, 2012;



Matthews and Lester, 2015) or based on performance indicators (e.g. liveweight gain in sheep and cattle; Kenyon and Jackson, 2012). The framework can be adapted to incorporate environmental stochasticity, as demonstrated here, to simulate the heterogeneity of intensity of infection between hosts that forms the basis of selection for TST. This was demonstrated by incorporating stochastic L3 intake rate and immune response in the present study.

Limitations of previously published detailed mechanistic approaches include an incomplete understanding of the processes and pathways involved in host immunity to GI nematode infection (although this is disputed by some; Roberts, 1999) and the detailed and invasive immunological datasets required to parameterise these models. The latter is a severe impediment to applying these models to understudied systems and necessitates a more simplified approach to modelling acquired immunity, regardless of whether or not there is an adequate understanding of the underlying processes. Complete representation of relevant immunological pathways, supported by sufficient empirical data to estimate parameters, is therefore difficult to achieve for most GI nematode species and is acknowledged as a bottleneck in the production of mathematical models for the population dynamics of GI nematodes (Charlier et al., 2018).

Previous models of GI nematode population dynamics and transmission have differed in their approaches to modelling acquired immunity, which increases during the course of an infection (Claerebout and Vercruysse, 2000). For example, some model acquired immunity as a simple increasing function of the time of exposure to infection (Learmount et al., 2006; Berk et al., 2016a), exposure to larvae (Grenfell et al., 1995), host age (Garnier et al., 2016), or worm burden (Cornell et al., 2004), or

mechanistically via the impact of exposure to infection on immunological parameters (Singleton et al., 2011; Prada Jiménez de Cisneros et al., 2014). Practical limitations of the former examples include the absence of upper boundaries on the levels of acquired immunity, which subsequently introduces difficulties scaling immune-mediated parasite life-history parameters. Practical limitations of the latter examples involve the need for invasive immunological measurements (plasma IgA) to estimate immune response rates.

The model presented here represents immunity as a logistic growth function. This allows an exponential response with cumulative exposure to GI nematodes, mimicking the stronger immune response that would be expected after administration of a challenge infection or a booster vaccination, and limits acquired immunity to values less than 1 to facilitate modelling interactions between host immune response and parasite life-history parameters. This simple approach also facilitates model application in data-sparse systems. To effectively model the development of host acquired immunity to GI nematodes without explicit representation of the complexities of the immune response and the necessary data for parameterisation, the decay and response rates for the logistic function are estimated using a combination of empirical, non-invasive field data, qualitative observational data and expert judgement. This approach requires fewer data for parameterisation than a more detailed mechanistic representation of immunity, and therefore permits application of this framework to a wider range of host-parasite systems than would be possible with a more detailed model. Nevertheless, there is scope to adapt the representation of immunity as more data become available. It would also be possible, with slight adaptation of the model, to apply varying levels of

host immunity to the different within-host life cycle stages, for example to simulate the use of vaccines with parasite stage-specific activity.

Overall, the mean-field model captured a high proportion of the variability in the observed FECs (*O. ostertagi* mean  $R^2 = 0.76$  and *C. oncophora* mean  $R^2 = 0.67$ ). Residual error can be attributed to multiple sources, including measurement error in the pasture larval counts used to initiate the model simulations (Couvillion, 1993), multiple sources of variability in the FEC method used in the validation dataset (Levecke et al., 2015), and individual host variability as described above. There was no evidence of systematic bias for *O. ostertagi*, and although the *C. oncophora* model tended to underestimate FECs, the differences in observed and predicted FECs were of no practical significance (section 2.4.4; COWS, 2014). The statistical significance of small deviations in predicted FECs from observed FECs (<25 eggs per gram in herds 1, 2, and 4; Fig. 3) highlights the importance of pragmatic validation methods including statistical, negative binomial models and qualitative appraisal. This is again highlighted by herd 7 which performed poorly in the statistical validation, despite low residual error and high  $R^2$  values (Table 3), and the predicted FECs being within a reasonable range of the observed FECs for both species tested. This was likely due to failure of the model to capture a slight, practically insignificant, increase in FEC for both species mid-August. Despite the overall good performance of the model, there were some exceptions. Herd 1 produced a high *C. oncophora* FEC 1 month after turnout which was not predicted by the model and could be of practical significance, as FECs of the level observed would usually require treatment. The FECs observed for herd 2 were significantly higher than predicted for both species, with differences at the end of the grazing season in the order of several

hundred egg. One plausible hypothesis for this, given the good performance of the model for most other herds, is that the individuals on this farm were particularly susceptible to GI nematode infection, which could be due to a number of factors such as genetics, nutrition and coinfection (which cannot be interrogated within the current dataset). Alternatively, the underprediction of FECs on this farm may indicate model bias when applied to young cattle – the calves simulated in Herd 2 were the youngest of all the farms used for model validation (6 months, cf. 10-21 months). Further validation would be required before applying the model to simulate very young calves (<10 months of age), to determine if this discrepancy is due to the susceptibility of the calves on this farm (posited above), an overestimate of immune response in younger calves, or a non-linear relationship between acquired immunity and the parasitological parameters.

To conclude, a generic framework to simulate the parasitic phase of GI nematode infections is presented here and its flexibility is demonstrated by simulating *O. ostertagi* and *C. oncophora* infections. The model simulations replicated infection patterns of first season grazers for these nematode species. The lead authors have previously developed a complementary framework for the free-living stages of the GI nematode life cycle, which has been applied to several GI nematode species, and has also been successfully adapted to simulate the development and dispersal of cattle lungworm (*Dictyocaulus viviparus*; McCarthy, C.A., 2018. Predicting the unpredictable: the changing epidemiology of *Dictyocaulus viviparus* in Great Britain. PhD Thesis. University of Liverpool, UK). It is hoped that GLOWORM-PARA will drive similar innovation and international collaboration by providing an accessible and transparent framework that can be adapted to multiple species and extended where

additional detail is required. Future research will integrate GLOWORM-PARA with GLOWORM-FL and host-parasite interactions (host movements, anthelmintic treatments etc.) to obtain a full lifecycle framework for the evaluation of alternative GI nematode control strategies.

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## Figure legends

Fig. 1. Conceptual framework of the GLOWORM-PARA model. State variable and parameter definitions are given in Table 1. Solid arrows indicate life-cycle transitions (e.g. from ingested L3 (L3i) to pre-adult (P) to adult (A)), mortality ( $\mu_i$ ) or deposition of eggs ( $\lambda$ ). Dashed arrows indicate dependencies (e.g. the level of acquired immunity ( $r$ ) depends on the intake of L3 (L3i)).

Fig. 2. Observed and predicted faecal egg counts (FEC) for *Ostertagia ostertagi* in first season grazing animals of dairy herds in Belgium. FECs were monitored for the entire length of the first grazing season. Further information on the background of this data can be found in Supplementary Table S1. Points and error bars show the observed number of eggs per gram of faeces (epg) and the corresponding 95% confidence interval obtained by bootstrapping (10,000 repeats). The dashed black line depicts the predicted FEC for a group of hosts based on a deterministic model simulation. The solid grey lines depict predictions obtained from 50 model simulations representing individual hosts, in which stochastic L3 intake and between-host variability in immune response were incorporated.

Fig. 3. Observed and predicted faecal egg counts (FEC) for *Cooperia oncophora* in first season grazing animals of dairy herds in Belgium. FECs were monitored for the entire length of the first grazing season. Further information on the background of this data can be found in Supplementary Table S1. Points and error bars show the observed number of eggs per gram of faeces (epg) and the corresponding 95%

confidence interval obtained by bootstrapping (10,000 repeats). The dashed black line depicts the predicted FEC for a group of hosts based on a deterministic model simulation. The solid grey lines depict predictions obtained from 50 model simulations representing individual hosts, in which stochastic L3 intake and between-host variability in immune response were incorporated.

Fig. 4. The observed and simulated faecal egg counts (FECs; points) with 95% confidence intervals (observed = horizontal, simulated = vertical). The diagonal black line indicates hypothetical perfect agreement between the observed and simulated FECs. The grey solid line indicates the predicted slope of the regression, with 95% confidence intervals show as grey dashed lines. Note that the 95% confidence intervals for the simulated data (estimated using the stochastic simulations shown in grey in Figs. 2 and 3) are narrow and may not be easily seen due to the scale of the Y-axes.

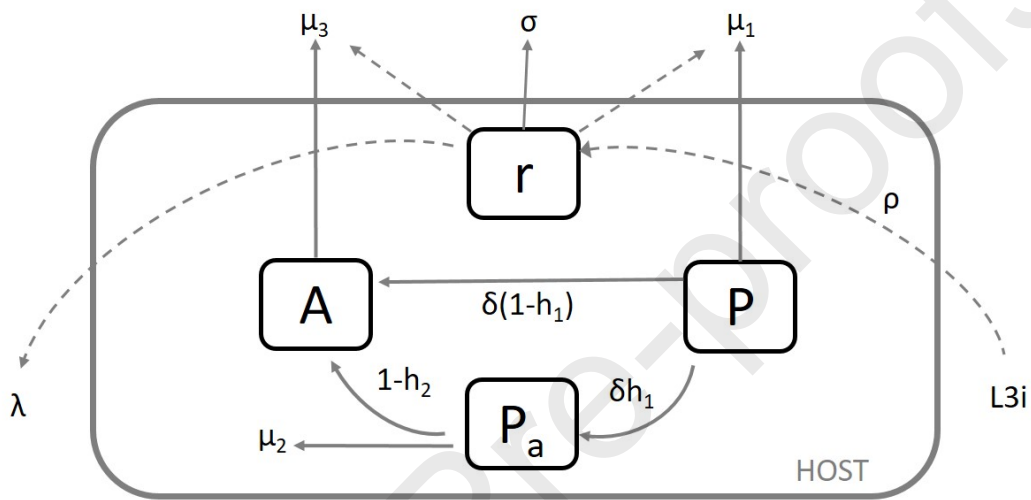
Supplementary Fig. S1. Mean daily temperature (A) and total daily rainfall (B) for the observation period for each herd used for model validation. Data were obtained from the E-OBS gridded dataset (Haylock et al., 2008) based on the village where each herd was located (Supplementary Table S1).

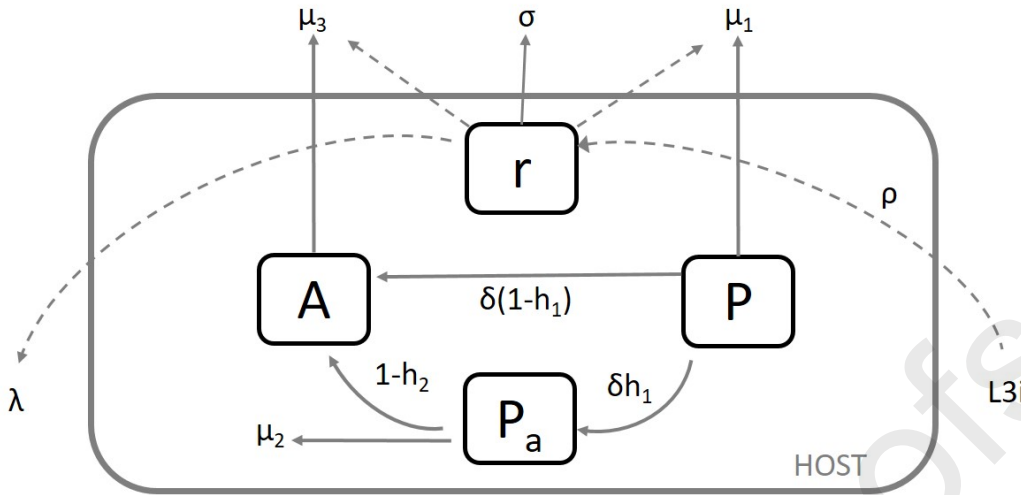
## Reference

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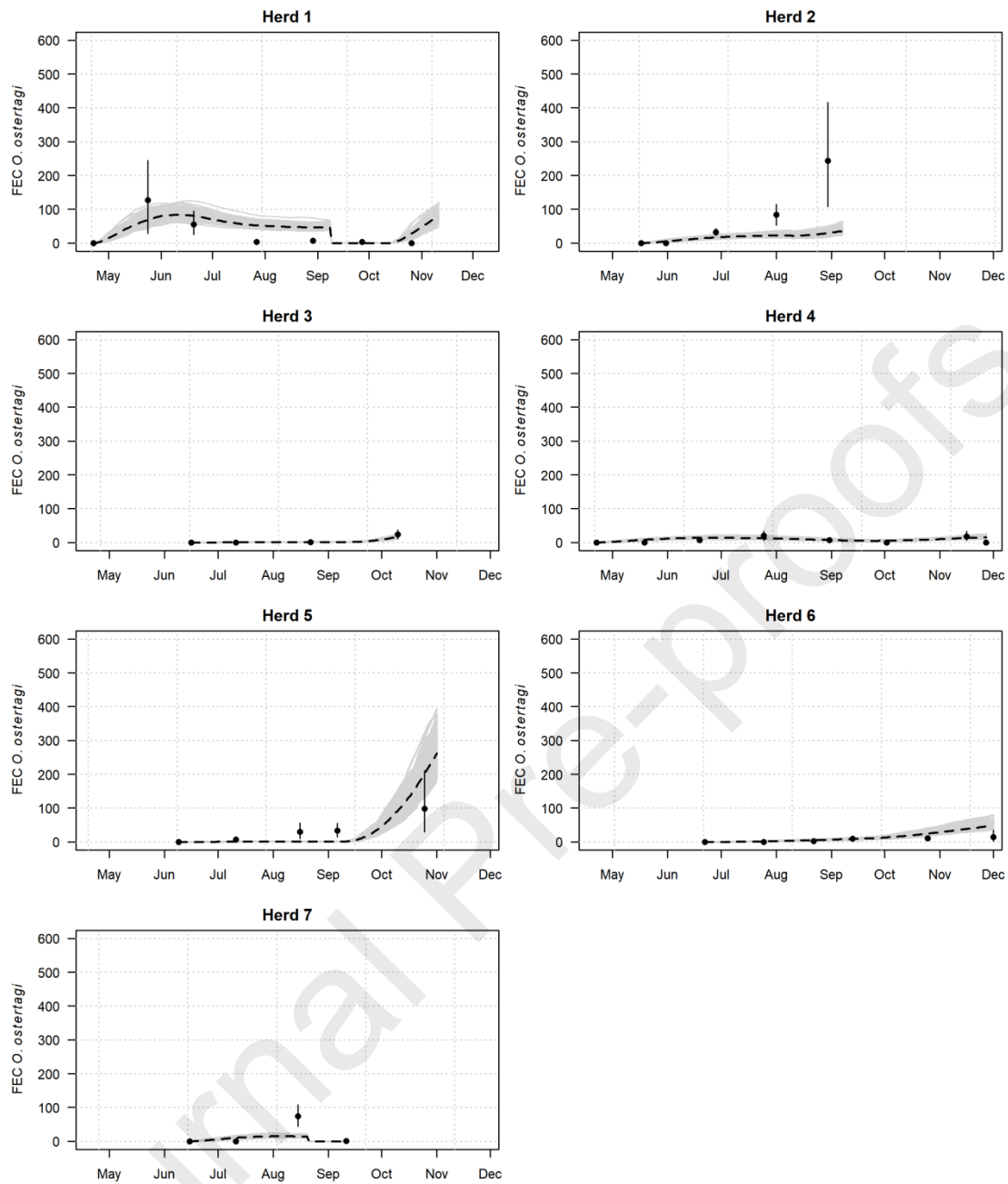
## Highlights

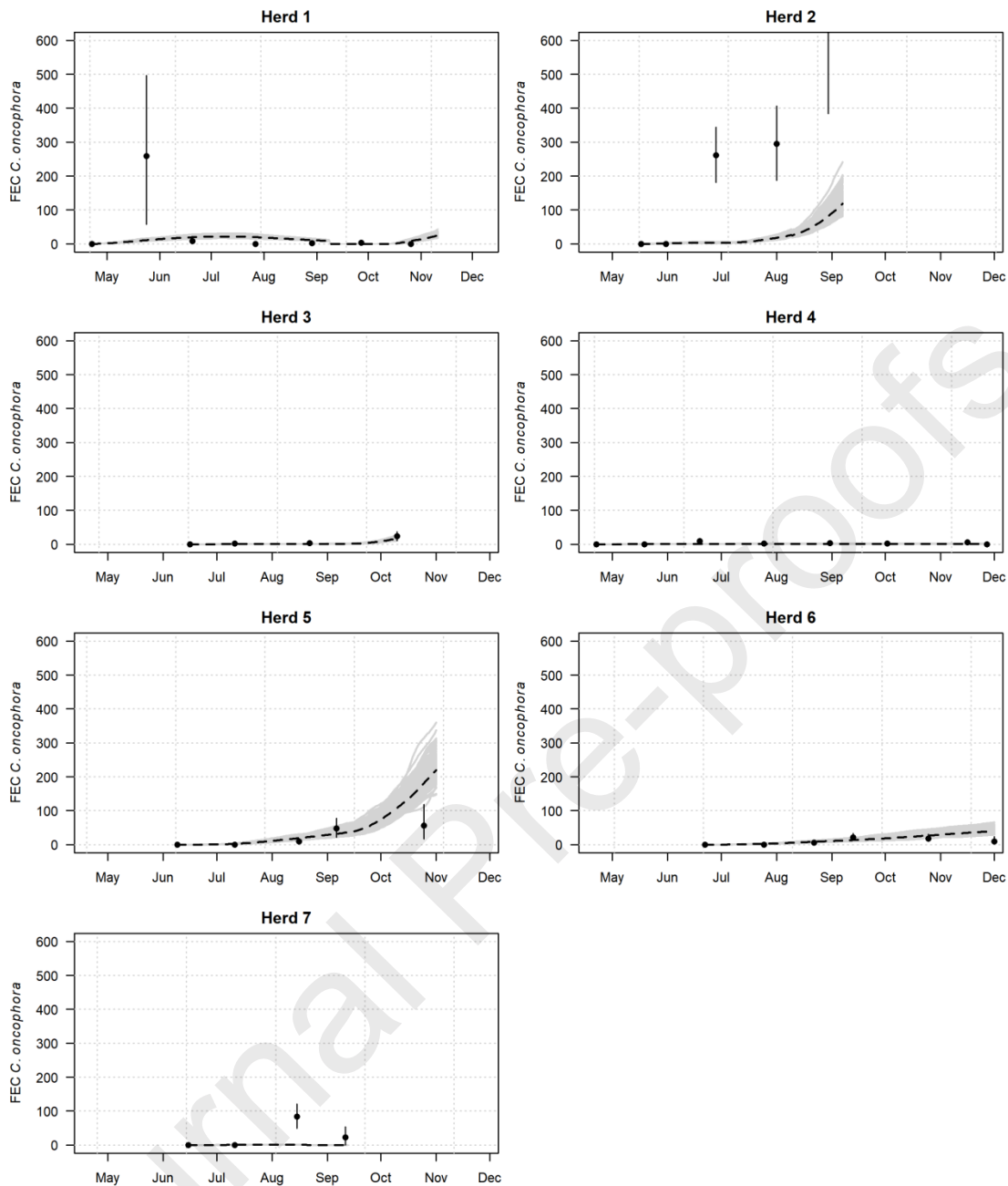
- The transmission of gastrointestinal nematode parasitic stages was modelled
- The generic model was parameterised for *Cooperia oncophora* and *Ostertagia ostertagi* in cattle
- Extensive validation using field data demonstrated good model performance
- A pragmatic approach to modelling data-sparse systems (immunity) is demonstrated
- Stochastic parameters can be introduced to incorporate host variability



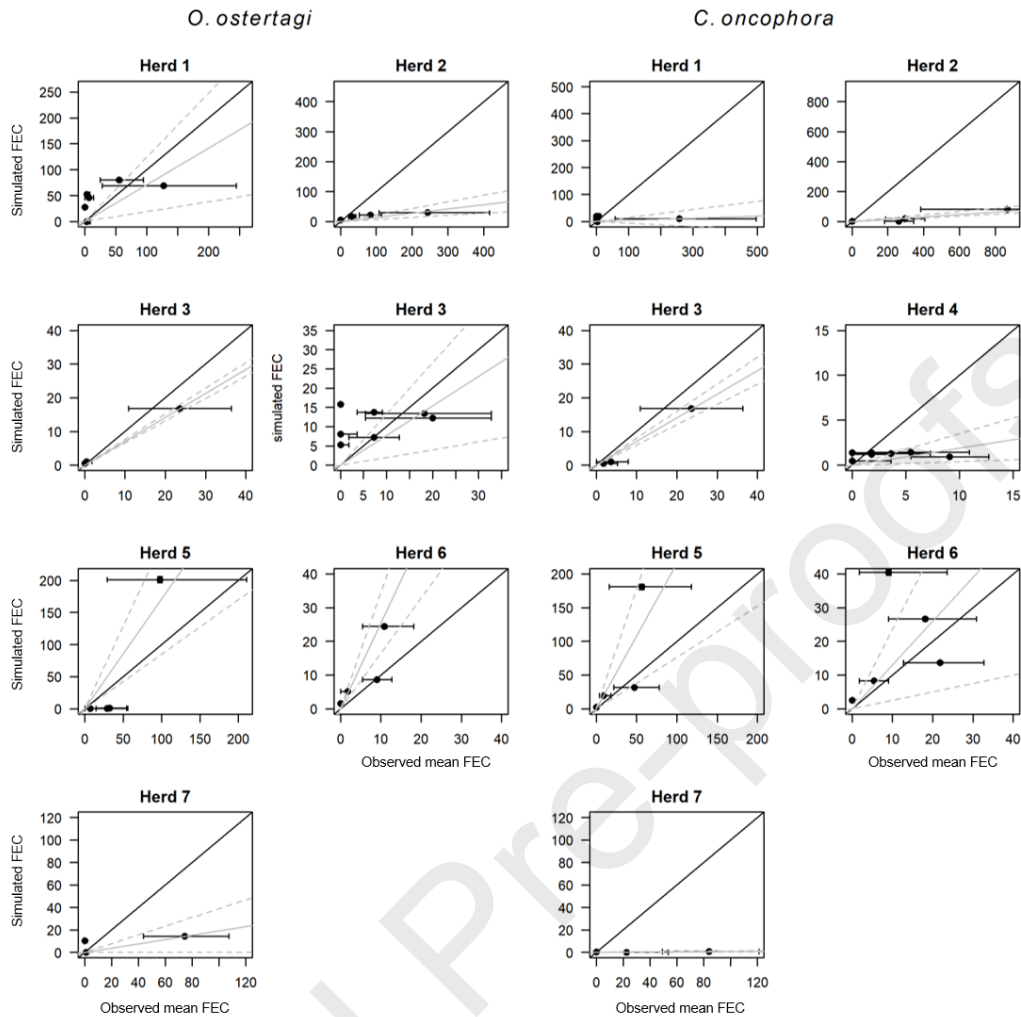


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951 **Table 1.** Model state variable and parameter definitions

State variable (bold) or Parameter	Definition	Units
<b><i>P</i></b>	Pre-adult nematodes in the host (L3, L4 and immature adult)	-
<b><i>Pa</i></b>	Arrested L4	-
<b><i>A</i></b>	Mature adult	-
<b><i>r</i></b>	Acquired immunity	-
<i>L3h</i>	L3 density on herbage	L3 kg dry matter <sup>-1</sup>
<i>L3i</i>	L3 ingestion rate	L3 day <sup>-1</sup> host <sup>-1</sup>
$\delta$	Development rate from ingested L3 to mature adult	P <sup>-1</sup> day <sup>-1</sup>
$\mu_1$	Pre-adult mortality rate	P <sup>-1</sup> day <sup>-1</sup>

$\mu_2$	Arrested L4 mortality rate	$\text{Pa}^{-1} \text{ day}^{-1}$
$\mu_3$	Adult mortality rate	$\text{A}^{-1} \text{ day}^{-1}$
$h_1$	Rate of developing pre-adult nematodes entering arrested development	$\text{P}^{-1} \text{ day}^{-1}$
$h_2$	Rate of arrested larvae resuming development	$\text{Pa}^{-1} \text{ day}^{-1}$
$\rho$	Immune response	$\text{L3i}^{-1}$
$\sigma$	Immune decay	$\text{Day}^{-1}$
$\lambda$	Daily fecundity (eggs produced)	$\text{Eggs worm}^{-1} \text{ day}^{-1}$
$f$	Daily faeces production	$\text{Grams (wet weight) day}^{-1}$
$d$	Driver of arrest	$\text{N/A}^a$
$\text{DMI}$	Daily herbage dry matter intake	$\text{Grams day}^{-1}$
$\text{kgDM}$	Standing biomass (herbage) on pasture	$\text{kg dry matter hectare}^{-1}$
$\text{FEC}$	Faecal egg count	$\text{Eggs gram}^{-1}$

<sup>a</sup> As this parameter is used to scale the immune-dependent parameters, the unit need not be fixed. In the present study, the temperature-dependent development rate of *Ostertagia ostertagi* or *Cooperia oncophora* has been used as an indicator of development success. However other, more complex, indicators can be used where sufficient data exist for parameterisation, such as  $Q_0$  estimates or the proportion of eggs surviving to L3 on pasture, or this parameter can be adapted to incorporate immunity-driven arrested development.

**Table 2.** Parameter estimates for *Ostertagia ostertagi* (*Oo*) and *Cooperia oncophora* (*Co*).

Parameter	Species	Estimate	Source
$\delta$	<i>Oo, Co</i>	$-\ln(0.5)/17 = 0.041$	Powers et al. (1982)
$\mu_{1(\min)}$	<i>Oo</i>	0.054	Mean in Verschave et al. (2014a)
	<i>Co</i>	0.044	Mean in Verschave et al. (2016b)
$\mu_{1(\max)}$	<i>Oo</i>	0.062	Upper 95% CI in Verschave et al. (2014a)
	<i>Co</i>	0.052	Upper 95% CI in Verschave et al. (2016b)
$\mu_2$	<i>Oo, Co</i>	0.002	Grenfell et al. (1987)
$\mu_{3(\min)}$	<i>Oo</i>	0.028	Mean in Verschave et al. (2014a)
	<i>Co</i>	0.039	Mean in Verschave et al. (2016b)
$\mu_{3(\max)}$	<i>Oo</i>	0.032	Upper 95% CI in Verschave et al. (2014a)

	<i>Co</i>	0.048	Upper 95% CI in Verschave et al. (2016b)
$h_{(min)}$	<i>Oo</i>	0.02	Lower 95% CI in Verschave et al. (2014a)
	<i>Co</i>	0.004	Lower 95% CI in Verschave et al. (2016b)
$h_{(max)}$	<i>Oo</i>	0.06	Upper 95% CI in Verschave et al. (2014a)
	<i>Co</i>	0.011	Upper 95% CI in Verschave et al. (2016b)
$\rho$	<i>Oo</i>	$5.981 \times 10^{-5}$	Current study (fitted to data from Shaw et al. (1998), see main text for assumptions)
	<i>Co</i>	$1.316 \times 10^{-4}$	Current study (fitted to data from Shaw et al. (1998), see section 2.3.3. for assumptions)
$\sigma$	<i>Oo, Co</i>	$-\ln(0.7)/(6 \times 30) = 0.002$	Current study (expert opinion)
$\lambda_{(min)}$	<i>Oo</i>	$\ln(196/2) = 4.58$	Lower 95% CI in Verschave et al. (2014a)
	<i>Co</i>	$\ln(1253/2) = 6.44$	Lower 95% CI in Verschave et al. (2016b)
$\lambda_{(max)}$	<i>Oo</i>	$\ln(284/2) = 4.96$	Mean in Verschave et al. (2014a), assuming a 1:1 sex ratio
	<i>Co</i>	$\ln(2968/2) = 7.30$	Mean in Verschave et al. (2016b); assuming a 1:1 sex ratio
$d$	<i>Oo</i>	$-0.07258 + 0.00976T^a$	Rose et al. (2015)
	<i>Co</i>	$-0.0401 + 0.00821T^a$	Current study (fitted to data from Sauermann and Leathwick (2018))

<sup>a</sup>T = daily mean air temperature (°C)

CI, Confidence Interval.

**Table 3.** Validation of simulations for faecal egg counts (FECs) of *Ostertagia ostertagi* and *Cooperia oncophora* using parasitological data of first season grazing animals on seven commercial dairy herds in Belgium.

Data set	<i>Ostertagia ostertagi</i>				<i>Cooperia oncophora</i>			
	Error (residual sum of squares)	Linear regression	$R^2$ ( $R^2$ adjusted)	Slope (95% CI)	Error (residual sum of squares)	Linear regression	$R^2$ ( $R^2$ adjusted)	Slope (95% CI)
<b>Her d 1</b>	37.69	$F_{1,5}=6.89$ , $P=0.047$	0.58 (0.50)	0.71 (0.19 – 1.24)	13.75	$F_{1,5}=0.69$ , $P=0.445$	0.12 (-0.006)	0.04 (-0.06 – 0.15)
<b>Her d 2</b>	10.11	$F_{1,3}=13.35$ , $P=0.035$	0.82 (0.76)	0.14 (0.07 – 0.22)	12.62	$F_{1,3}=41.06$ , $P=0.008$	0.93 (0.91)	0.08 (0.06 – 0.11)
<b>Her d 3</b>	0.61	$F_{1,2}=766.7$ , $P=0.001$	1 (1)	0.71 (0.66 – 0.76)	1.21	$F_{1,2}=191.3$ , $P=0.005$	0.99 (0.98)	0.70 (0.60 – 0.80)
<b>Her d 4</b>	8.38	$F_{1,6}=7.00$ , $P=0.038$	0.54 (0.46)	0.77 (0.20 – 1.33)	0.92	$F_{1,6}=5.90$ , $P=0.051$	0.50 (0.41)	0.19 (0.04 – 0.35)
<b>Her d 5</b>	47.31	$F_{1,3}=15.1$ , $P=0.030$	0.83 (0.78)	1.71 (0.85 – 2.57)	53.02	$F_{1,3}=9.12$ , $P=0.057$	0.75 (0.67)	2.16 (0.76 – 3.56)
<b>Her d 6</b>	9.42	$F_{1,4}=30.35$ , $P=0.005$	0.88 (0.85)	2.54 (1.64 – 3.45)	16.25	$F_{1,4}=5.9$ , $P=0.07$	0.60 (0.50)	1.30 (0.25 – 2.35)
<b>Her d 7</b>	7.39	$F_{1,2}=3.84$ , $P=0.189$	0.66 (0.49)	0.19 (-0.000 – 0.39)	0.24	$F_{1,2}=6.48$ , $P=0.126$	0.76 (0.65)	0.007 (0.00 – 0.013)

CI, Confidence Interval.